Questions about the class...

- Registration/waitlist questions:
 - gradinfo@stat.ubc.ca
- Labs start this Thursday.
- Lab registration (about 10 students):

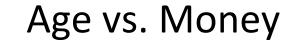
 Send an email with what labs are available to you: gradinfo@stat.ubc.ca
- Midterm; Lecture Slides; Webwork
 - webwork.elearning.ubc.ca

Questions about the class...

- Midterm; Lecture Slides; Webwork...
- Pre-requisite knowledge:
 - One sample and two sample t-tests
 - Hypothesis tests
 - Confidence Intervals
 - Type 1 error and Type 2 error
 - probability density function (pdf), cumulative density function (cdf).
 - Properties of a Normal distribuiton.

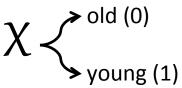
Stat 306: Finding Relationships in Data. Lecture 2 Least Squares for one predictor

t-test

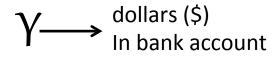




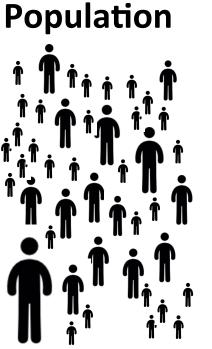
Dependent variable



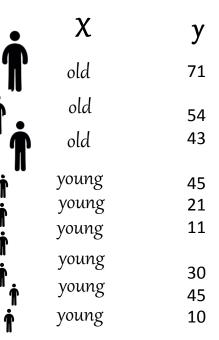




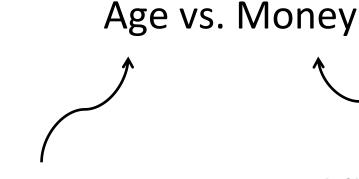
Sample, n=9



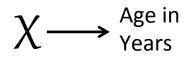
Population parameters μ_0, μ_1, σ^2 Hypothesis Test $H_0: \mu_0 = \mu_1$ $H_1: \mu_0 \neq \mu_1$ Sample statistics $\bar{y}_0 = 56$ $\bar{y}_1 = 27$ $\bar{y}_0 - \bar{y}_1 = 29$ $s_p = 10.81$ t = 2.68, df = 7p-value = 0.03 95% C.I. = [3.4, 54.6]



linear regression

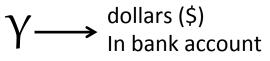


PREDICTOR variable





RESPONSE variable



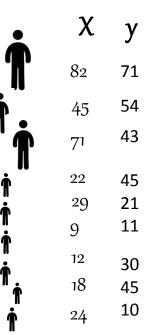
Sample, n=9

Population $\dot{\mathbf{n}} \quad \dot{\mathbf{n}} \quad \dot{$

Population parameters β_0 , β_1 , σ^2

Hypothesis Test $H_0: \beta_1 = 0$ $H_1: \beta_1 \neq 0$ Sample statistics $b_0 = 17.7$ $b_1 = 0.55$ s = 15.5 $R^2 = 0.49$

For statistic b_1 : 95% C.I. = [0.05, 1.05] *p*-value = 0.036



Chapter 2

Simple linear regression

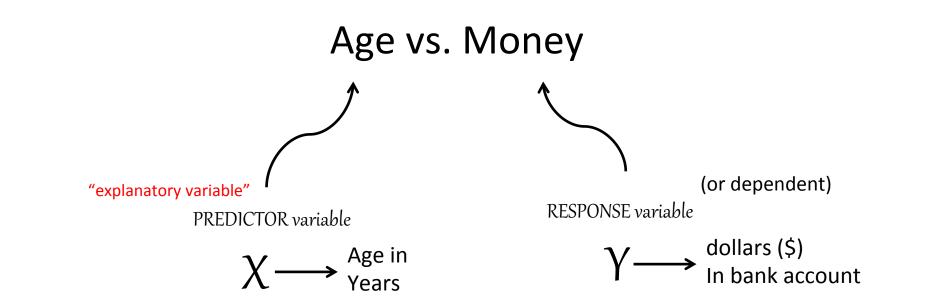
In forming prediction equations, there are typically several explanatory variables x_1, \ldots, x_p that affect the response variable y. Given data vectors $(x_{i1}, \ldots, x_{ip}, y_i)$ for $i = 1, \ldots, n$, the goal is the find a prediction equation that goes through the middle of the data points.

This chapter starts with the study of the simple case of one explanatory variable (p = 1). The data vectors are written as (x_i, y_i) for i = 1, ..., n and the simplest prediction equation to consider is a straight line that goes through the middle of the scatterplot of y versus x.

The theory of estimation of the prediction equation with least squares involves calculus, and the statistical inference for predictions involve a statistical or probability model for the random deviations from a prediction equation. The usual assumption (for the first attempted prediction equation) is that deviations from a theoretical prediction equation are independent normal random variables with a mean of 0 and a variance that does not depend on the values of the explanatory variables.

For one explanatory variable, derivations of estimators and interval estimates can be done without matrices, and it is easier to show the meaning of different quantities via scatterplots. The same steps with matrices and vectors will be carried out in Chapter 3 when there is more than one explanatory variable.

- We will start with the simple case of one **explanatory variable**.
- The data vectors are written as (x_i, y_i) for i = 1, ..., n.
- The simplest **prediction equation** to consider is a straight line that goes through the middle of the scatterplot of y versus x.

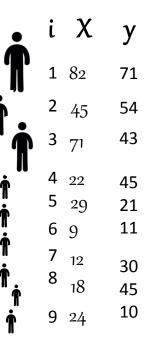


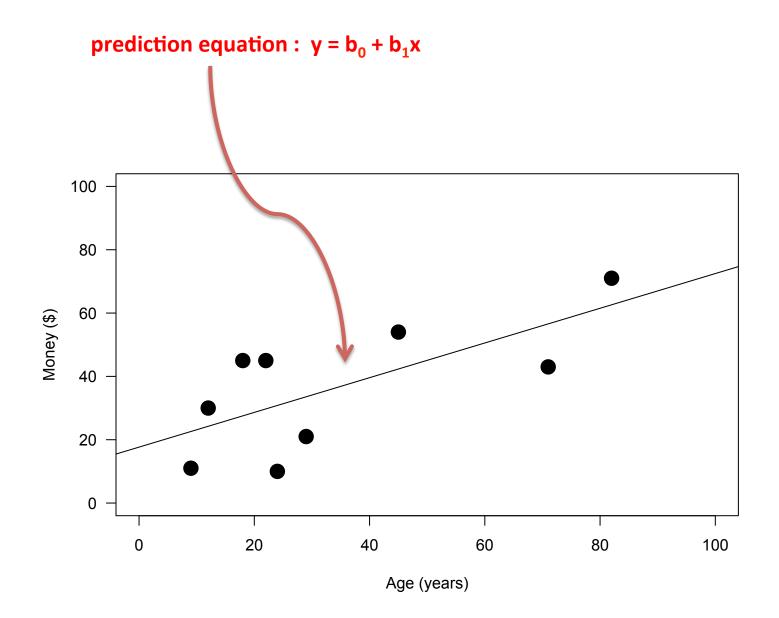
data vectors

Population parameters β_0 , β_1 , σ^2

Hypothesis Test $H_0: \beta_1 = 0$ $H_1: \beta_1 \neq 0$ Sample statistics $b_0 = 17.7$ $b_1 = 0.55$ s = 15.5 $R^2 = 0.49$

For parameter β_1 : 95% C.I. = [0.05, 1.05] *p*-value = 0.036





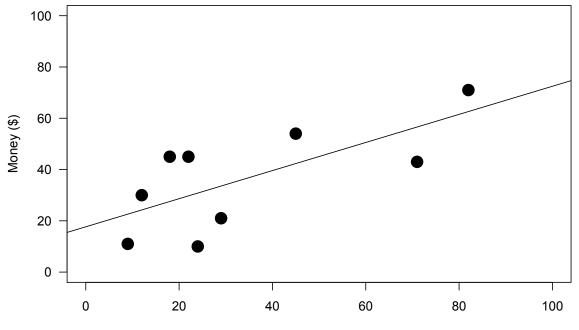
- We will start with the simple case of one **explanatory variable**.
- The data vectors are written as (xi , yi) for i = 1, . . . , n.
- The simplest **prediction equation** to consider is a straight line that goes through the middle of the scatterplot of y versus x.

Chapter 2

- Section 2.1 has the mathematics leading to the least squares line.
- Section 2.2 introduces the simple linear regression model (prediction with one explanatory variable) that is formulated for a predictive equation. This is needed to quantify the variability of the coefficients of the best-fitting line, when different samples are taken from the population.
- Section 2.5 has intervals for simple linear regression: the confidence interval for the slope of the least square line, confidence intervals for subpopulation means, and prediction intervals for a future or out-of-sample Y given x*.
- Section 2.6 has an explanation of Student t quantiles used in the interval estimates.

Section 2.1

"finding a best-fitting line to a scatterplot of a Y variable versus an X variable"



Age (years)

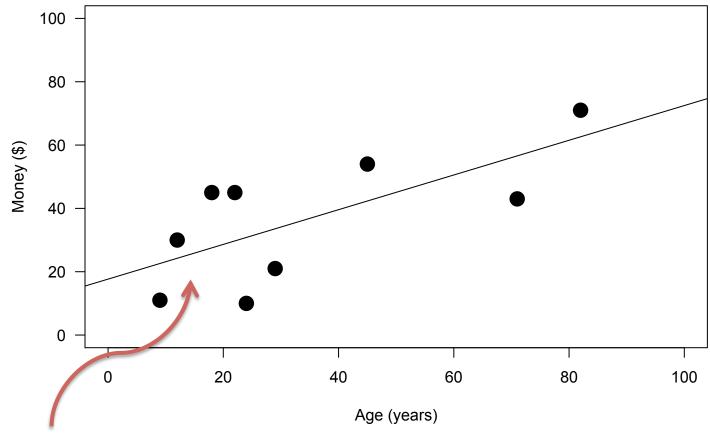
Section 2.1

"finding a best-fitting line to a scatterplot of a Y variable versus an X variable"

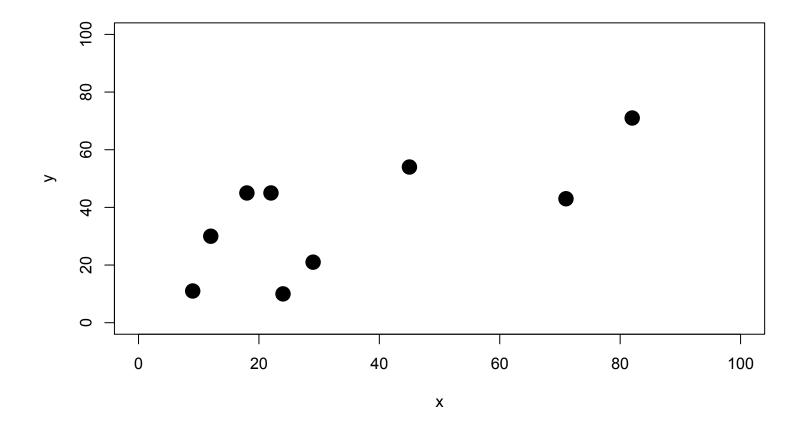
• 2.1.1 Visualization exercises

• 2.1.2 Summary statistics

• 2.1.3 Least squares solution



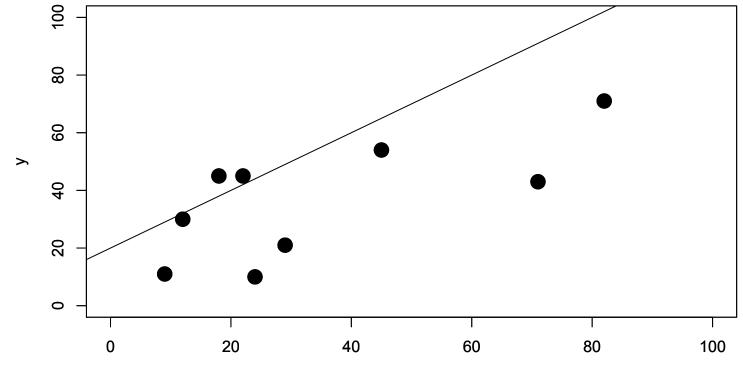
prediction equation



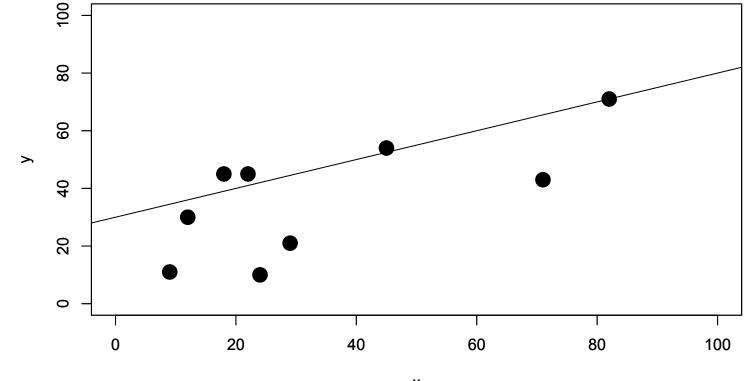
> x <- c(82, 45, 71, 22, 29, 9, 12, 18, 24)
> y <- c(71, 54, 43, 45, 21, 11, 30, 45, 10)</pre>

> n <- 9

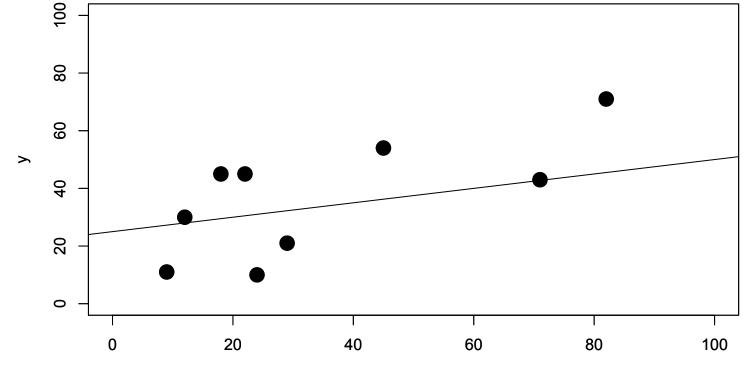
> plot(y~x, pch=20, cex=3, xlim=c(0,100), ylim=c(0,100))



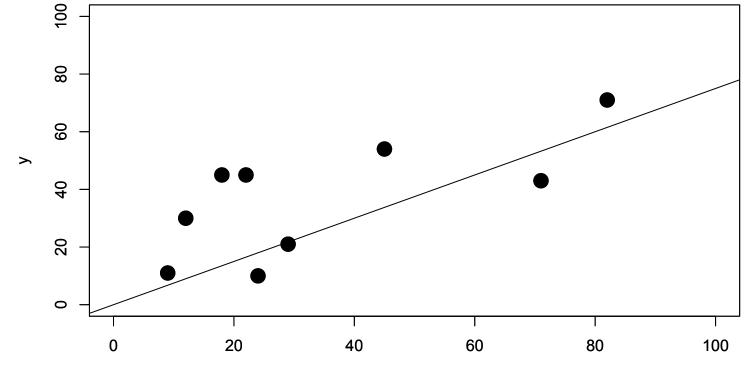
- > x <- c(82, 45, 71, 22, 29, 9, 12, 18, 24)
- > y <- c(71, 54, 43, 45, 21, 11, 30, 45, 10)
- > n <- 9
- > plot(y~x, pch=20, cex=3, xlim=c(0,100), ylim=c(0,100))
- > abline(20,1)



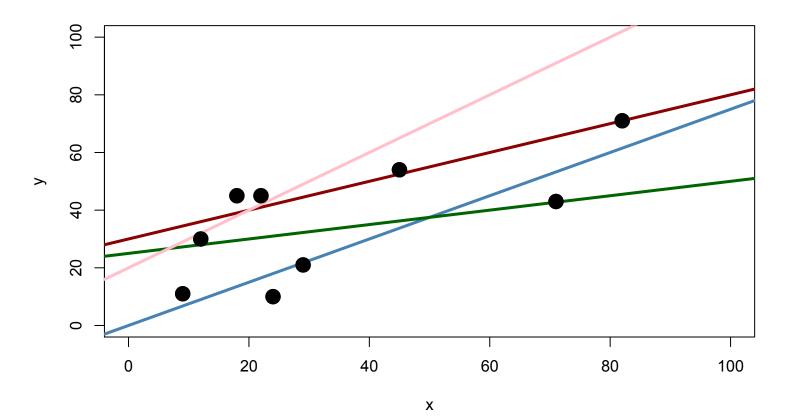
- > x <- c(82, 45, 71, 22, 29, 9, 12, 18, 24)
- > y <- c(71, 54, 43, 45, 21, 11, 30, 45, 10)
- > n <- 9
- > plot(y~x, pch=20, cex=3, xlim=c(0,100), ylim=c(0,100))
- > abline(30,0.5)



- > x <- c(82, 45, 71, 22, 29, 9, 12, 18, 24)
- > y <- c(71, 54, 43, 45, 21, 11, 30, 45, 10)
- > n <- 9
- > plot(y~x, pch=20, cex=3, xlim=c(0,100), ylim=c(0,100))
- > abline(25,0.25)



- > x <- c(82, 45, 71, 22, 29, 9, 12, 18, 24)
- > y <- c(71, 54, 43, 45, 21, 11, 30, 45, 10)
- > n <- 9
- > plot(y~x, pch=20, cex=3, xlim=c(0,100), ylim=c(0,100))
- > abline(0,0.75)



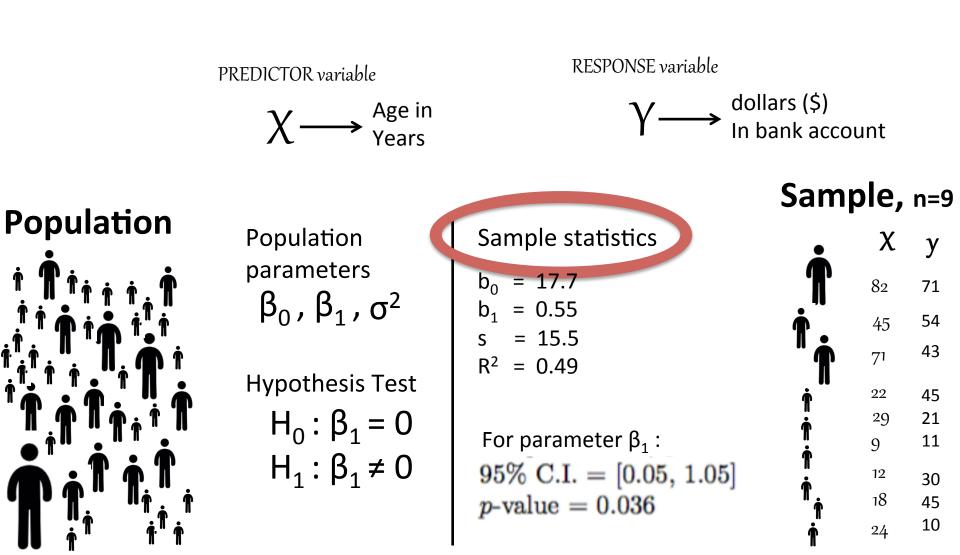
Which of the four lines (visually) looks like the best-fitting line?

y = 0 + 0.75xy = 25 + 0.25x y = 30 + 0.5x y = 20 + 1x

Guess the regression line:

<u>https://www.geogebra.org/m/JsFmFEg6</u>

<u>https://www.geogebra.org/m/B7JtA6Mg</u>



The data are first summarized into sample means (measures of centre)

$$ar{x} = n^{-1}\sum_{i=1}^n x_i$$

 $ar{y} = n^{-1}\sum_{i=1}^n y_i$

> xbar<-(1/n)*sum(x)
> xbar
[1] 34.666667

> ybar<-(1/n)*sum(y)
> ybar
[1] 36.66667

The data are first summarized into sample means (measures of centre) and standard deviations (measures of spread):

$$\overline{x} = n^{-1} \sum_{i=1}^{n} x_i, \quad s_x = \sqrt{\sum_{i=1}^{n} (x_i - \overline{x})^2 / (n-1)},$$

$$\overline{y} = n^{-1} \sum_{i=1}^{n} y_i, \quad s_y = \sqrt{\sum_{i=1}^{n} (y_i - \overline{y})^2 / (n-1)}.$$

> xbar<-(1/n)*sum(x) > ybar<-(1/n)*sum(y) > xbar [1] 34.66667 [1] 36.66667 > > sx<-sqrt(sum((x-xbar)^2)/(n-1)) > sy<-sqrt(sum((y-ybar)^2)/(n-1)) > sx [1] 26.03843 [1] 20.36541

Correlation describes the linear association between X and Y.

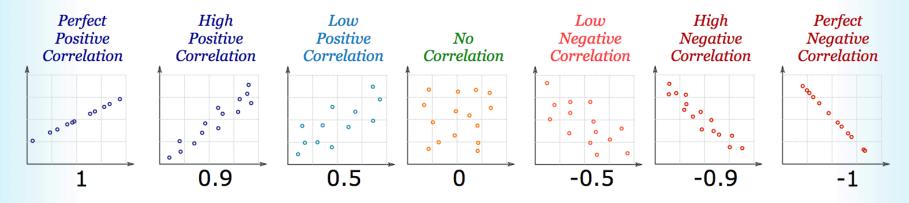
- Correlation is **Positive** when the values **increase** together, and
- Correlation is Negative when one value decreases as the other increases

https://www.mathsisfun.com/data/correlation.html

Guess the Correlation Game: http://guessthecorrelation.com/

- Correlation is **Positive** when the values **increase** together, and
- Correlation is Negative when one value decreases as the other increases

Here we look at **linear correlations** (correlations that follow a line).



Correlation can have a value:

- 1 is a perfect positive correlation
- 0 is no correlation (the values don't seem linked at all)
- -1 is a perfect negative correlation

https://www.mathsisfun.com/data/correlation.html

Guess the Correlation Game: http://guessthecorrelation.com/

To summarize the linear association, the sample correlation is

$$r_{xy} = \frac{s_{xy}}{s_x s_y}$$

where sample covariance is

$$s_{xy} = (n-1)^{-1} \sum_{i=1}^n (x_i - \overline{x})(y_i - \overline{y}).$$

To summarize the linear association, the sample correlation is $r_{xy} = \frac{s_{xy}}{s_x s_y}$, where sample covariance is

(2.3)
$$s_{xy} = (n-1)^{-1} \sum_{i=1}^{n} (x_i - \overline{x})(y_i - \overline{y}).$$

sample covariance:

To summarize the linear association, the sample correlation is $r_{xy} = \frac{s_{xy}}{s_x s_y}$, where sample covariance is

(2.3)
$$s_{xy} = (n-1)^{-1} \sum_{i=1}^{n} (x_i - \overline{x})(y_i - \overline{y}).$$

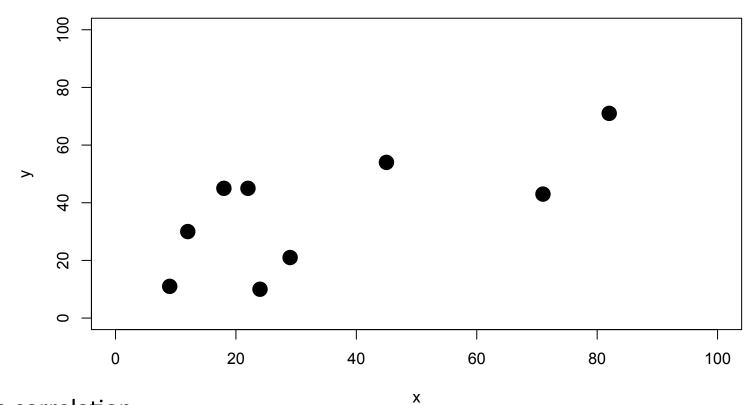
sample covariance:

```
> sxy<-(1/(n-1))*sum((x-xbar)*(y-ybar))
> sxy
[1] 371.625
```

Sample correlation:

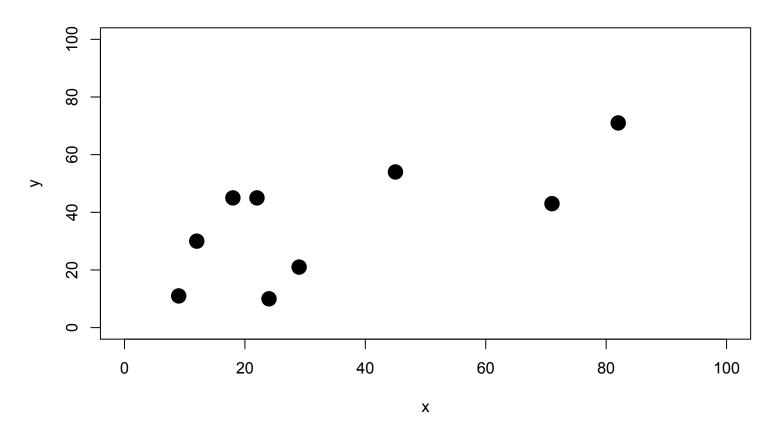
```
> rxy<-sxy/(sx*sy)
> rxy
[1] 0.7008045
```

Note that $-1 \leq r_{xy} \leq 1$ is a scaled version of the unbounded s_{xy} , that is, $-\infty < s_{xy} < \infty$. The quantities r_{xy} and s_{xy} are positive (negative) if the scatterplot of (x_i, y_i) , i = 1, ..., n, form a cloud the slopes upwards (downwards).



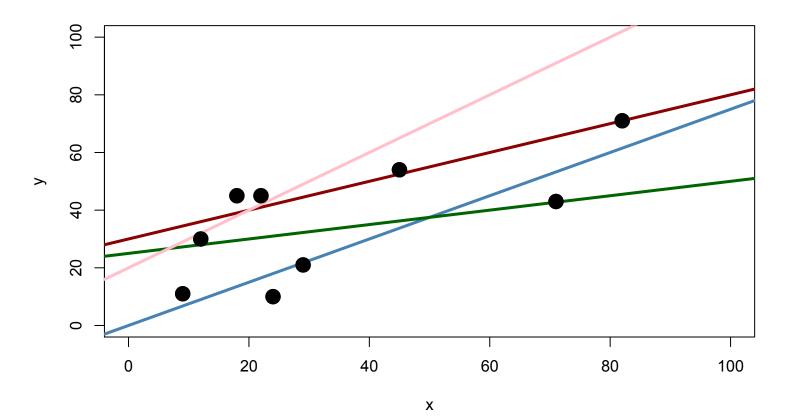
Sample correlation:

> rxy<-sxy/(sx*sy)
> rxy
[1] 0.7008045



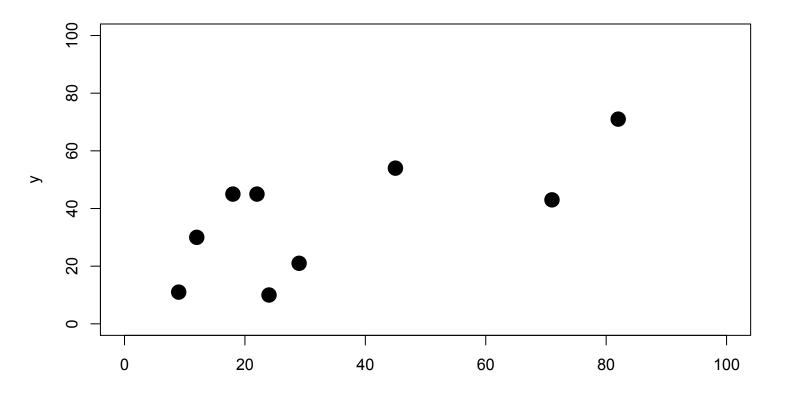
Data don't lie on a straight line or smooth curve because there is noise or variation from other factors that affect the response.

If the two variables are linearly related, a summary is a line $y = b_0 + b_1 x$ that goes through the middle of the scatterplot.



Which of the four lines (visually) looks like the best-fitting line?

y = 0 + 0.75xy = 25 + 0.25x y = 30 + 0.5x y = 20 + 1x

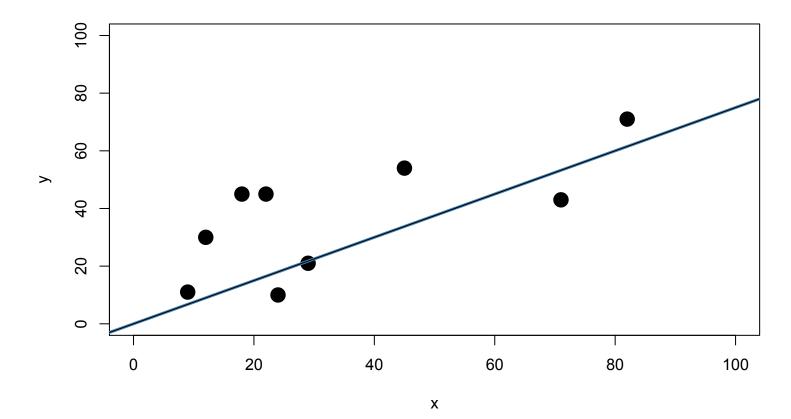


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To find b₀ and b₁ of a best-fitting line, one choice is to minimize **the Sum of Squared Errors**:

$$S(b_0,b_1) = \sum_{i=1}^n (y_i - b_0 - b_1 x_i)^2$$

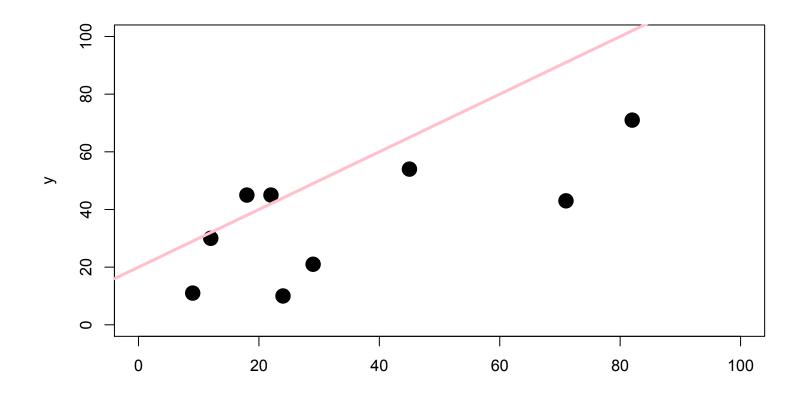
What is the Sum of Squared Errors? <u>https://www.geogebra.org/m/JsFmFEg6</u> or <u>http://students.brown.edu/seeing-theory/regression/index.html</u>



To find b_0 and b_1 of a best-fitting line, one choice is to minimize the **SSE**:

$$S(b_0, b_1) = \sum_{i=1}^{n} (y_i - b_0 - b_1 x_i)^2$$

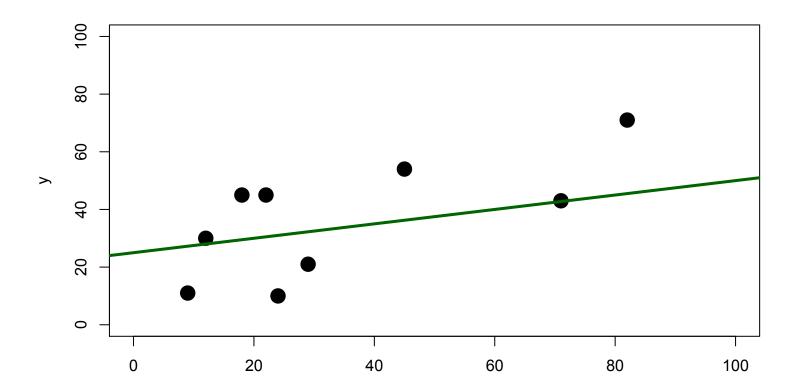
> b0<-0; b1<-0.75
> sum((y- b0 - b1*x)^2)
[1] 2933.5



To find b_0 and b_1 of a best-fitting line, one choice is to minimize:

$$S(b_0, b_1) = \sum_{i=1}^{n} (y_i - b_0 - b_1 x_i)^2$$

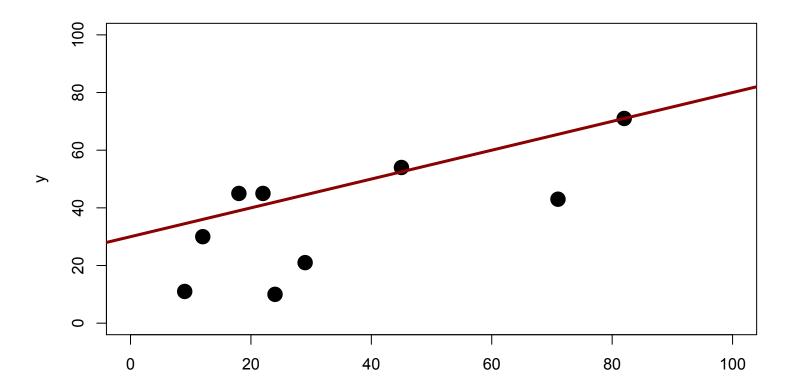
> b0<-20; b1<-1
> sum((y- b0 - b1*x)^2)
[1] 5712



To find b_0 and b_1 of a best-fitting line, one choice is to minimize:

$$S(b_0, b_1) = \sum_{i=1}^{n} (y_i - b_0 - b_1 x_i)^2$$

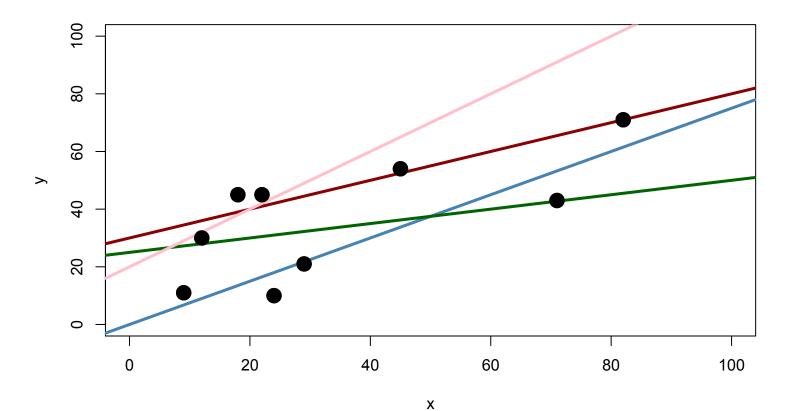
> b0<-25; b1<-0.25
> sum((y- b0 - b1*x)^2)
[1] 2251.5



To find b_0 and b_1 of a best-fitting line, one choice is to minimize:

$$S(b_0, b_1) = \sum_{i=1}^{n} (y_i - b_0 - b_1 x_i)^2$$

> b0<-30; b1<-0.5
> sum((y- b0 - b1*x)^2)
[1] 2725



Which of the four lines (visually) looks like the best-fitting line?

y = 0 + 0.75x	S(b0,b1) = 2933.5
y = 25 + 0.25x	S(b0,b1) = 2251.5
y = 30 + 0.5x	S(b0,b1) = 2725.0
y = 20 + 1x	S(b0,b1) = 5712.0

To find b_0 and b_1 of a best-fitting line, one choice is to minimize:

$$S(b_0,b_1) = \sum_{i=1}^n (y_i - b_0 - b_1 x_i)^2$$

The line that minimizes $S(b_0, b_1)$ is called the:

least squares regression line (LSRL)

To find b_0 and b_1 of a best-fitting line, one choice is to minimize:

$$S(b_0,b_1) = \sum_{i=1}^n (y_i - b_0 - b_1 x_i)^2$$

The line that minimizes $S(b_0, b_1)$ is called the:

least squares regression line (LSRL)

It minimizes the sum of squared errors (also known as: "SSE" or "squared vertical deviations" or "Sum of Squares").

Question: How do we find the values of b_0 and b_1 that minimize $S(b_0, b_1)$?

Answer: Simple calculus

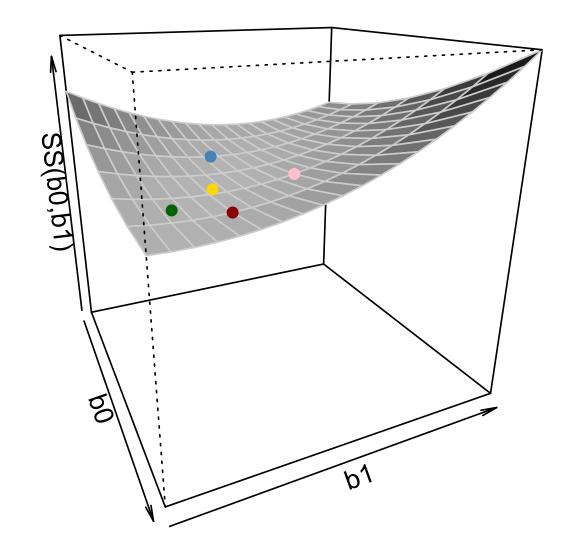
$$S(b_0,b_1) = \sum_{i=1}^n (y_i - b_0 - b_1 x_i)^2$$

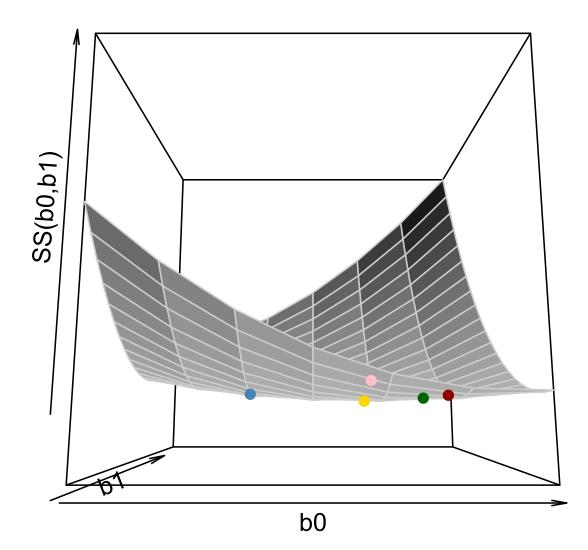
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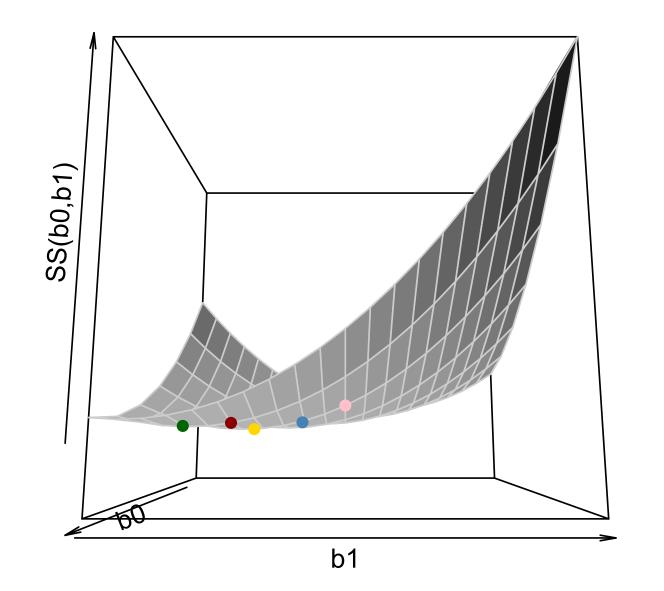
The goal is to minimize $S(b_0, b_1) = \sum_{i=1}^{n} (y_i - b_0 - b_1 x_i)^2$.

The partial derivatives of S are:

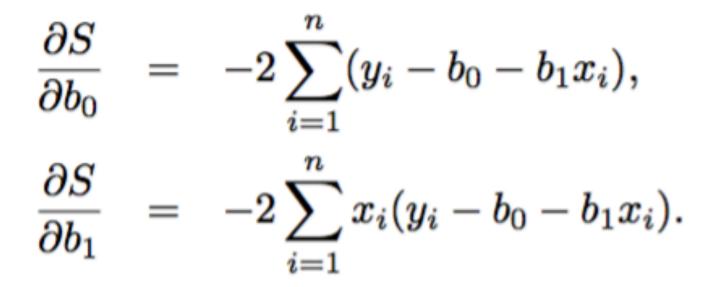
$$egin{array}{rcl} \displaystylerac{\partial S}{\partial b_0}&=&-2\sum_{i=1}^n(y_i-b_0-b_1x_i),\ \displaystylerac{\partial S}{\partial b_1}&=&-2\sum_{i=1}^nx_i(y_i-b_0-b_1x_i) \end{array}$$







The goal is to minimize $S(b_0, b_1) = \sum_{i=1}^{n} (y_i - b_0 - b_1 x_i)^2$.



Set the equations to 0, divide by -2 and solve.

TIP: Go through this process with a pen and paper until you can easily obtain the solution yourself.

The goal is to minimize $S(b_0, b_1) = \sum_{i=1}^{n} (y_i - b_0 - b_1 x_i)^2$.

Set the equations to 0, divide by -2 and solve.

The solution (\hat{b}_0, \hat{b}_1) satisfies

$$0 = n[\overline{y} - \hat{b}_0 - \hat{b}_1 \overline{x}],$$

$$0 = \sum_{i=1}^n x_i y_i - \hat{b}_0 n \overline{x} - \hat{b}_1 \sum_{i=1}^n x_i^2.$$

$$0 = n[\overline{y} - \hat{b}_0 - \hat{b}_1 \overline{x}],$$

$$0 = \sum_{i=1}^n x_i y_i - \hat{b}_0 n \overline{x} - \hat{b}_1 \sum_{i=1}^n x_i^2.$$

$$(2.18) \qquad \qquad \hat{b}_0 = \overline{y} - \hat{b}_1 \overline{x},$$

$$0 = n[\overline{y} - \hat{b}_0 - \hat{b}_1 \overline{x}],$$

$$0 = \sum_{i=1}^n x_i y_i - \hat{b}_0 n \overline{x} - \hat{b}_1 \sum_{i=1}^n x_i^2.$$

(2.18)
$$\hat{b}_0 = \overline{y} - \hat{b}_1 \overline{x},$$

(2.19)
$$0 = \sum_{i=1}^n x_i y_i - [\overline{y} - \hat{b}_1 \overline{x}] n \overline{x} - \hat{b}_1 \sum_{i=1}^n x_i^2,$$

$$0 = n[\overline{y} - \hat{b}_0 - \hat{b}_1 \overline{x}],$$

$$0 = \sum_{i=1}^n x_i y_i - \hat{b}_0 n \overline{x} - \hat{b}_1 \sum_{i=1}^n x_i^2.$$

(2.18)
$$\hat{b}_{0} = \overline{y} - \hat{b}_{1}\overline{x},$$

(2.19) $0 = \sum_{i=1}^{n} x_{i}y_{i} - [\overline{y} - \hat{b}_{1}\overline{x}]n\overline{x} - \hat{b}_{1}\sum_{i=1}^{n} x_{i}^{2},$
(2.20) $0 = \sum_{i=1}^{n} x_{i}y_{i} - n\overline{x}\overline{y} + n\hat{b}_{1}\overline{x}^{2} - \hat{b}_{1}\sum_{i=1}^{n} x_{i}^{2},$

(2.18)
$$\hat{b}_{0} = \overline{y} - \hat{b}_{1}\overline{x},$$

(2.19) $0 = \sum_{i=1}^{n} x_{i}y_{i} - [\overline{y} - \hat{b}_{1}\overline{x}]n\overline{x} - \hat{b}_{1}\sum_{i=1}^{n} x_{i}^{2},$
(2.20) $0 = \sum_{i=1}^{n} x_{i}y_{i} - n\overline{x}\overline{y} + n\hat{b}_{1}\overline{x}^{2} - \hat{b}_{1}\sum_{i=1}^{n} x_{i}^{2},$
(2.21) $\hat{b}_{1} = \frac{\sum_{i=1}^{n} x_{i}y_{i} - n\overline{x}\overline{y}}{\sum_{i=1}^{n} x_{i}^{2} - n\overline{x}^{2}}$

$$(2.18) \qquad \hat{b}_{0} = \overline{y} - \hat{b}_{1}\overline{x},$$

$$(2.19) \qquad 0 = \sum_{i=1}^{n} x_{i}y_{i} - [\overline{y} - \hat{b}_{1}\overline{x}]n\overline{x} - \hat{b}_{1}\sum_{i=1}^{n} x_{i}^{2},$$

$$(2.20) \qquad 0 = \sum_{i=1}^{n} x_{i}y_{i} - n\overline{x}\overline{y} + n\hat{b}_{1}\overline{x}^{2} - \hat{b}_{1}\sum_{i=1}^{n} x_{i}^{2},$$

$$(2.21) \qquad \hat{b}_{1} = \frac{\sum_{i=1}^{n} x_{i}y_{i} - n\overline{x}\overline{y}}{\sum_{i=1}^{n} x_{i}^{2} - n\overline{x}^{2}}$$

$$(2.22) \qquad = \frac{(n-1)s_{xy}}{(n-1)s_{x}^{2}}$$

$$(2.18) \qquad \hat{b}_{0} = \overline{y} - \hat{b}_{1}\overline{x},$$

$$(2.19) \qquad 0 = \sum_{i=1}^{n} x_{i}y_{i} - [\overline{y} - \hat{b}_{1}\overline{x}]n\overline{x} - \hat{b}_{1}\sum_{i=1}^{n} x_{i}^{2},$$

$$(2.20) \qquad 0 = \sum_{i=1}^{n} x_{i}y_{i} - n\overline{x}\overline{y} + n\hat{b}_{1}\overline{x}^{2} - \hat{b}_{1}\sum_{i=1}^{n} x_{i}^{2},$$

$$(2.21) \qquad \hat{b}_{1} = \frac{\sum_{i=1}^{n} x_{i}y_{i} - n\overline{x}\overline{y}}{\sum_{i=1}^{n} x_{i}^{2} - n\overline{x}^{2}}$$

$$(2.22) \qquad = \frac{(n-1)s_{xy}}{(n-1)s_{x}^{2}}$$

$$(2.23) \qquad = \frac{r_{xy}s_{x}s_{y}}{s_{x}^{2}} = \frac{r_{xy}s_{y}}{s_{x}}.$$

 $\hat{b}_0 \;\;=\;\; \overline{y} - \hat{b}_1 \overline{x},$ (2.18)

- -

(2.19)
$$0 = \sum_{i=1}^{n} x_i y_i - [\overline{y} - \hat{b}_1 \overline{x}] n \overline{x} - \hat{b}_1 \sum_{i=1}^{n} x_i^2,$$

n

(2.20)
$$0 = \sum_{i=1}^{n} x_i y_i - n \overline{xy} + n \hat{b}_1 \overline{x}^2 - \hat{b}_1 \sum_{i=1}^{n} x_i^2,$$

(2.21)
$$\hat{b}_1 = \frac{\sum_{i=1}^n x_i y_i - n\overline{xy}}{\sum_{i=1}^n x_i^2 - n\overline{x}^2}$$

(2.22)
$$= \frac{(n-1)s_{xy}}{(n-1)s_x^2}$$

(2.23)
$$= \frac{r_{xy}s_xs_y}{s_x^2} = \frac{r_{xy}s_y}{s_x}.$$

The solution is therefore:

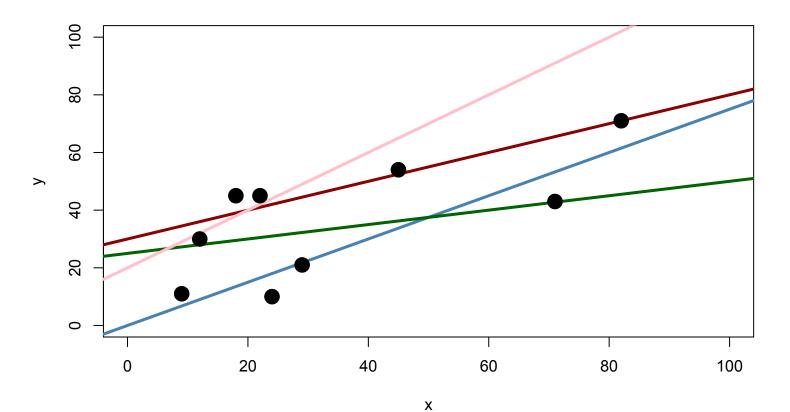
$$\hat{b}_0 = \overline{y} - \hat{b}_1 \overline{x}$$

 $\hat{b}_1 - r + \hat{s}_1 \langle s \rangle$

 $v_1 - r_{xy}s_{y/s_x}$

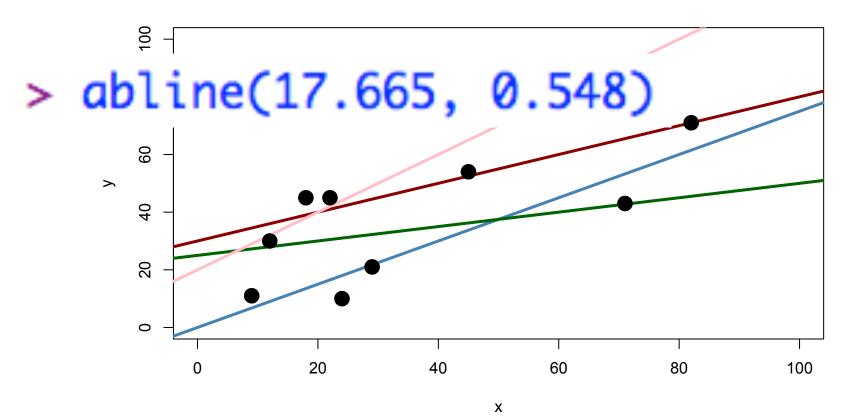
```
2.1.3 Least squares solution
                                               \hat{b}_0 = \overline{y} - \hat{b}_1 \overline{x}
  The solution is therefore:
                                               \hat{b}_1 = r_{xy} s_y / s_x
> xbar < -(1/n) * sum(x)
> xbar
[1] 34.66667
> sx<-sqrt( sum((x-xbar)^2)/(n-1) )</pre>
> SX
[1] 26.03843
>
                                                 > rxy<-sxy/(sx*sy)</pre>
> ybar<-(1/n)*sum(y)</pre>
> ybar
                                                 > rxy
                                                  [1] 0.7008045
[1] 36.66667
> sy<-sqrt( sum((y-ybar)^2)/(n-1) )</pre>
                                                 > b1_hat<-rxy*sy/sx</pre>
                                                 > b0_hat<-ybar-b1_hat*xbar</p>
> SY
[1] 20.36541
                                                  >
                                                 > b1_hat
                                                  [1] 0.5481195
> sxy<-(1/(n-1))*sum((x-xbar)*(y-ybar))</pre>
                                                 > b0_hat
> SXY
                                                  [1] 17.66519
[1] 371.625
```

2.1.3 Least squares solution $\hat{b}_0 = \overline{y} - \hat{b}_1 \overline{x}$ The solution is therefore: $\hat{b}_1 = r_{xy} s_y / s_x$ > b1_hat<-rxy*sy/sx</p> > b0_hat<-ybar-b1_hat*xbar</pre> > > b1_hat [1] 0.5481195 > b0_hat [1] 17.66519 > > b0<-17.665; b1<-0.548 $> sum((y-b0 - b1*x)^2)$ [1] 1688.441



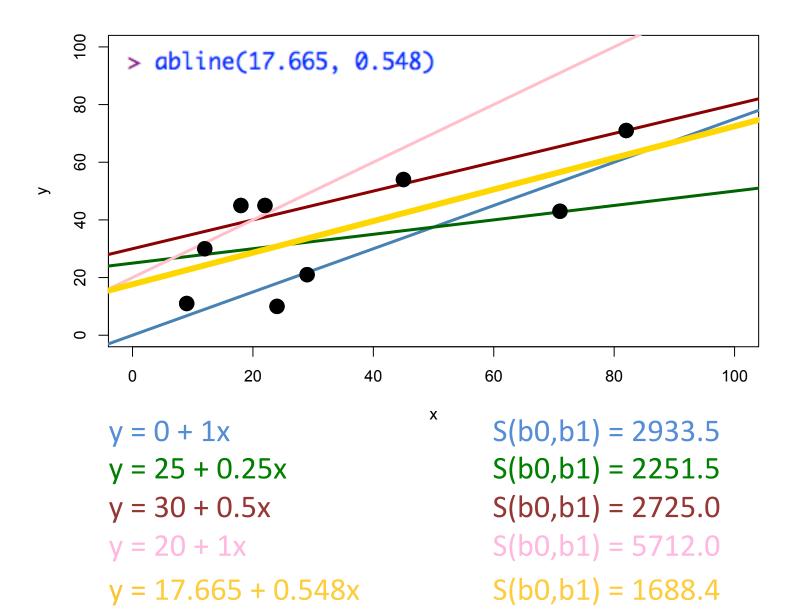
Which of the four lines (visually) looks like the best-fitting line?

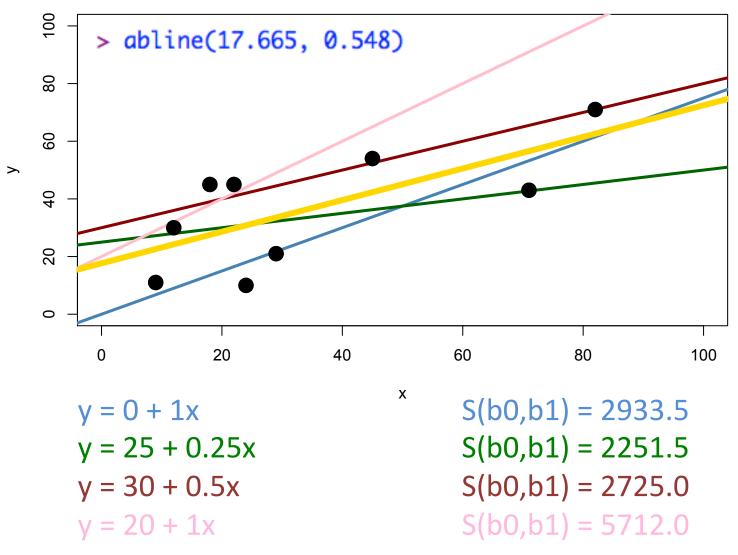
y = 0 + 1x	S(b0,b1) = 2933.5
y = 25 + 0.25x	S(b0,b1) = 2251.5
y = 30 + 0.5x	S(b0,b1) = 2725.0
y = 20 + 1x	S(b0,b1) = 5712.0



Which of the four lines (visually) looks like the best-fitting line?

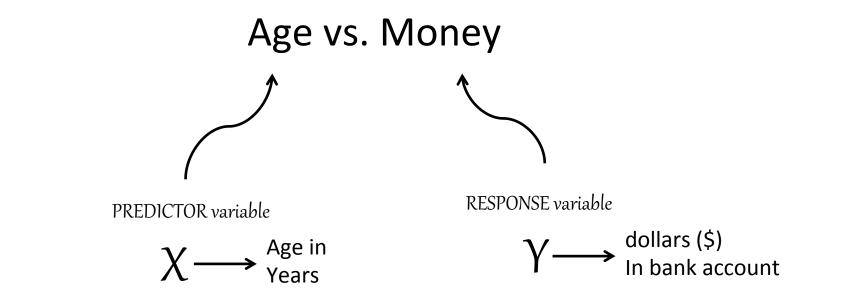
y = 0 + 1x	S(b0,b1) = 2933.5
y = 25 + 0.25x	S(b0,b1) = 2251.5
y = 30 + 0.5x	S(b0,b1) = 2725.0
y = 20 + 1x	S(b0,b1) = 5712.0



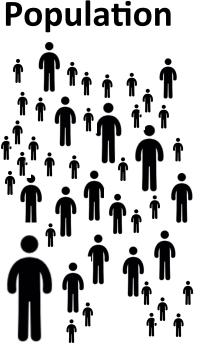


y = 17.7 + 0.55x

S(b0,b1) = 1688.4

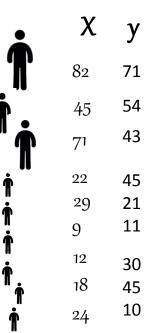


Sample, n=9



Population parameters β_0 , β_1 , σ^2

Hypothesis Test $H_0: \beta_1 = 0$ $H_1: \beta_1 \neq 0$ Sample statistics $b_0 = 17.7$ $b_1 = 0.55$ s = 15.5 $R^2 = 0.49$ For parameter β_1 : 95% C.I. = [0.05, 1.05] *p*-value = 0.036



Age vs. Money

Sample statistics

		17.7
b ₁	=	0.55
S	=	15.5
R ²	=	0.49

For statistic b_1 : 95% C.I. = [0.05, 1.05] *p*-value = 0.036

We obtained a random sample of n = 9 subjects. There is a statistically significant association between age and money (p-value =0.036).
 For every additional year in age, an individual's amount of money increases on average by an estimated of \$0.55 (95% C.I. = [\$0.05, \$1.05]).

We collected a random sample of individuals and for each

determined their age (recorded in years) and the amount

of money (in dollars) in their accounts. Analysis of

Conclusions: We found that, as hypothesized, age is associated with money. In our sample age accounted for about half of the variability observed in money (R²=0.49). We **predict** that a 50 year old will have \$45.1 (95% P.I. = [\$5.6, \$84.5]), whereas a 40 year old will have \$39.6 (95% P.I. = [\$0.8, \$78.4]).

The purpose of this observational study was to

demonstrate if, and to what extent, age is

the data was done using **linear regression**.

associated with money.

Small Print: The analysis rests on the following assumptions:

Objective:

Design and Methods:

- the observations are independently and identically distributed.
- the **response** variable, money, is normally distributed.
- Homoscedasticity of residuals or equal variance.
- the <u>relationship</u> between **response** and **predictor** variables is linear.

(2.24)
$$\hat{y} = \hat{b}_0 + \hat{b}_1 x = (\overline{y} - \hat{b}_1 \overline{x}) + \hat{b}_1 x$$

(2.24)
$$\hat{y} = \hat{b}_0 + \hat{b}_1 x = (\overline{y} - \hat{b}_1 \overline{x}) + \hat{b}_1 x$$
$$(2.25) = \overline{y} + \hat{b}_1 (x - \overline{x})$$

(2.24)

$$\begin{aligned}
\hat{y} &= \hat{b}_0 + \hat{b}_1 x = (\overline{y} - \hat{b}_1 \overline{x}) + \hat{b}_1 x \\
&= \overline{y} + \hat{b}_1 (x - \overline{x}) \\
&= \overline{y} + s_y r_{xy} \frac{x - \overline{x}}{s_x},
\end{aligned}$$

$$(2.24) \qquad \hat{y} = \hat{b}_0 + \hat{b}_1 x = (\overline{y} - \hat{b}_1 \overline{x}) + \hat{b}_1 x$$

$$(2.25) \qquad \qquad = \overline{y} + \hat{b}_1 (x - \overline{x})$$

$$(2.26) \qquad \qquad = \overline{y} + s_y r_{xy} \frac{x - \overline{x}}{s_x},$$

$$(2.27) \qquad \qquad \qquad \frac{\hat{y} - \overline{y}}{s_y} = r_{xy} \frac{x - \overline{x}}{s_x}.$$

In terms of summary statistics, the least squares line is:

Equation (2.27) is the easiest way to remember the least squares line (read as "standardized y = correlation coefficient times standardized x"); from this, the intercept \hat{b}_0 and slope \hat{b}_1 can be obtained using simple algebra.

Chapter 2

- Section 2.1 has the mathematics leading to the least squares line.
- Section 2.2 introduces the simple linear regression model (prediction with one explanatory variable) that is formulated for a predictive equation. This is needed to quantify the variability of the coefficients of the best-fitting line, when different samples are taken from the population.
- Section 2.5 has intervals for simple linear regression: the confidence interval for the slope of the least square line, confidence intervals for subpopulation means, and prediction intervals for a future or out-of-sample Y given x*.
- Section 2.6 has an explanation of Student t quantiles used in the interval estimates.